BMC Bioinformatics



Software Open Access

Cubic exact solutions for the estimation of pairwise haplotype frequencies: implications for linkage disequilibrium analyses and a web tool 'CubeX'

Tom R Gaunt*, Santiago Rodríguez and Ian NM Day

Address: Bristol Genetic Epidemology Laboratories (BGEL) and MRC Centre for Causal Analyses in Translational Epidemiology (CAITE), Department of Social Medicine, University of Bristol, Canynge Hall, Whiteladies Road, Bristol, BS8 2PR, UK

Email: Tom R Gaunt* - tom.gaunt@bristol.ac.uk; Santiago Rodríguez - santi.rodriguez@bristol.ac.uk; Ian NM Day - ian.day@bristol.ac.uk * Corresponding author

Published: 2 November 2007

BMC Bioinformatics 2007, 8:428 doi:10.1186/1471-2105-8-428

Received: 5 February 2007 Accepted: 2 November 2007

This article is available from: http://www.biomedcentral.com/1471-2105/8/428

© 2007 Gaunt et al; licensee BioMed Central Ltd.

This is an Open Access article distributed under the terms of the Creative Commons Attribution License (http://creativecommons.org/licenses/by/2.0), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

Abstract

Background: The frequency of a haplotype comprising one allele at each of two loci can be expressed as a cubic equation (the 'Hill equation'), the solution of which gives that frequency. Most haplotype and linkage disequilibrium analysis programs use iteration-based algorithms which substitute an estimate of haplotype frequency into the equation, producing a new estimate which is repeatedly fed back into the equation until the values converge to a maximum likelihood estimate (expectation-maximisation).

Results: We present a program, "CubeX", which calculates the biologically possible exact solution(s) and provides estimated haplotype frequencies, D', r^2 and χ^2 values for each. CubeX provides a "complete" analysis of haplotype frequencies and linkage disequilibrium for a pair of biallelic markers under situations where sampling variation and genotyping errors distort sample Hardy-Weinberg equilibrium, potentially causing more than one biologically possible solution. We also present an analysis of simulations and real data using the algebraically exact solution, which indicates that under perfect sample Hardy-Weinberg equilibrium there is only one biologically possible solution, but that under other conditions there may be more.

Conclusion: Our analyses demonstrate that lower allele frequencies, lower sample numbers, population stratification and a possible |D'| value of I are particularly susceptible to distortion of sample Hardy-Weinberg equilibrium, which has significant implications for calculation of linkage disequilibrium in small sample sizes (eg HapMap) and rarer alleles (eg paucimorphisms, q < 0.05) that may have particular disease relevance and require improved approaches for meaningful evaluation.

Background

Linkage disequilibrium (LD) describes the condition that occurs when alleles at different loci are non-randomly associated in a given population. Under LD the frequency (f_{11}) of a haplotype (h_{11}) representing the "1" allele at two

loci is significantly more or less than the product of the respective allele frequencies. Characterisation of LD is important in medical genetics, influencing association mapping of trait loci and providing information on interactions between genes [1,2]. LD is the result of a shared

history of mutation and recombination, and other factors including: genetic drift, population growth, admixture, population structure, the ages of the polymorphisms, the physical distance separating them and the effects of selective pressure [3].

For unrelated individuals the estimation of LD relies on the estimation of haplotype frequencies. In a 3 × 3 table for a biallelic marker the haplotype phase of all individuals is known with the exception of the centre cell (representing individuals heterozygous at both loci). The estimated frequency, \hat{f}_{11} , of the haplotype h_{11} is described by a cubic equation of the form

$$a\hat{f}_{11}^{3} + b\hat{f}_{11}^{2} + c\hat{f}_{11} + d = 0$$
 (1)

that is adapted from Hill's equation (4) [4] with the constants defined under Methods. With \hat{f}_{11} and the allele frequencies, all four haplotype frequencies can be calculated, thus estimating the unknown proportions of the middle cell.

Several approaches exist for solving equation (1), the solution of which enables estimation of haplotype frequencies and LD coefficients. The first approach uses iteration-based algorithms. An initial estimate of haplotype frequency (either random, or based on the known haplotype numbers) is substituted into the equation, providing a new estimate. This is then fed back into the equation and the expectation-maximisation (EM) process repeated until the values converge. This is the basis both of the algorithm described by Hill in 1974 for the estimation of pairwise haplotype frequencies [4], and of other EM algorithms that enable the estimation of multilocus haplotype frequencies. Many programs exist that utilise variations on this approach, including: GOLD [5], GOLDsurfer [6], MIDAS [7], Haploview [8] and many others reviewed in [9-12]. The potential problem for these approaches is that algorithms may converge on one of the alternative roots of the cubic equation (a local maximum rather than the global maximum).

Other approaches include parsimony, eg HAPAR [13] and Bayesian algorithms, eg PHASE [14-16]. Parsimony and Bayesian methods are both better suited to estimating individual haplotypes than EM approaches, while Bayesian and EM methods are useful for estimating population frequencies [11].

An alternative approach would be exact solution, such as *Cardan's solution* [17] of the generalized cubic equation (of which equation (1) is an example). This provides all roots to the cubic equation, from which we can select those that

are both *real* (i.e. not a complex number) and *biologically possible*. If more than one solution exists then the likelihoods of the different solutions can be compared and an informed evaluation made of the result. Theoretically, the non-iterative approach may be computationally less intensive and more accurate, but computational efficiency and accuracy will be software and platform dependent.

Implementation

Hill assumed random mating and Hardy Weinberg Equilibrium (HWE) [4]. Rearranging terms for consequent diplotype frequency expectations for two biallelic loci Luo and Suhai [18] obtained equation 1 given in the introduction (here redefining \hat{f}_{11} as x, a3 as a, a2 as b, a1 as c and a as d for convenience): $ax^3 + bx^2 + cx + d = 0$, where a = 4n; $b = 2n (1 - 2p - 2q) - 2(2n_{11} + n_{12} + n_{21}) - n_{22}; c = 2npq (2n_{11} + n_{12} + n_{21})(1 - 2p - 2q) - n_{22}(1 - p - q); d = -(2n_{11} + n_{22})(1 - p - q); d = -(2n_{11} + n_{12})(1 - p - q); d$ $n_{12} + n_{21}$)pq; n = number of subjects; p = common allele freq of locus 1; q = common allele freq of locus 2; n_{11} is the number of subjects who are homozygous for the commoner allele at both loci; n_{12} are common homozygous at locus 1 and heterozygous at locus 2; n_{21} are heterozygous at locus 1 and common homozygous at locus 2; n_{22} are heterozygous at both loci [18]. Equation 1 can be solved exactly for x (with 1 to 3 real number solutions).

We have adopted the Nickalls treatment of the Cardan solution of the generalized cubic equation [17], and written a Python [19] program "CubeX" to solve equation 1 exactly. In CubeX, after calculation of constants *a-d* from diplotypic data the following are calculated:

$$x_N = -b/(3a)$$
; $\delta^2 = (b^2 - 3ac)/9a^2$; $h^2 = 4a^2\delta^6$; $\gamma_N = ax_N^3 + bx_N^2 + cx_N + d$.

The discriminant $\Delta_3 = \gamma_N^2 - h^2$ is then used to determine the outcome in real roots (without having to go through complex number intermediates or ambiguities), with three possible outcomes:

Outcome 1: if $\gamma_N^2 > h^2$ there will be only one real root (α) given by

$$\alpha = x_N + \sqrt[3]{\frac{1}{2a} \left(-\gamma_N + \sqrt{\gamma_N^2 - h^2} \right)} + \sqrt[3]{\frac{1}{2a} \left(-\gamma_N - \sqrt{\gamma_N^2 - h^2} \right)}$$
(2)

Outcome 2: if $\gamma_N^2 = h^2$ there are three real roots (α , β and γ) and α and β are equal. For a value of $\mu = \sqrt[3]{\frac{\gamma_N}{2a}}$:

$$\alpha = x_N + \mu \tag{3}$$

$$\beta = x_N + \mu \tag{4}$$

$$\gamma = x_N - 2\mu \tag{5}$$

Outcome 3: if $\gamma_N^2 < h^2$ there are three real roots (α , β and γ). Where $\theta = \frac{\arccos(-\gamma_N/h)}{3}$:

$$\alpha = x_N + 2\delta \cos \theta \tag{6}$$

$$\beta = x_N + 2\delta \cos(2\pi/3 + \theta) \tag{7}$$

$$\gamma = x_N + 2\delta\cos(4\pi/3 + \theta) \tag{8}$$

Values for D' and r^2 are calculated as previously described [20,21]:

$$D = (f_{11} \times f_{22}) - (f_{12} \times f_{21})$$

 $D_{max} = min [p(1-q), (1-p)q] \text{ if } D > 0 \text{ or } D_{max} = min [pq, (1-p)(1-q)] \text{ if } D < 0$

$$D' = D/D_{max}$$
 (9)

$$r^2 = D^2/(p(1-p)q(1-q))$$
 (10)

Diplotype frequencies based on the estimated haplotype frequencies are compared to the input diplotype frequencies by a χ^2 test, which effectively tests sample deviation from the null hypothesis of HWE for the diplotypes formed of the four haplotypes. The number of degrees of freedom is equal to the number of observations (diplotype counts) minus four estimated parameters which are either three haplotypes (the fourth can be inferred) and D, or one haplotype, two allele frequencies and D. If nine different diplotypes are observed the number of degrees of freedom is therefore five. For each empty cell in the 3×3 the number of degrees of freedom is reduced by one. If the user knows there are only three haplotypes present (and therefore six diplotypes) then there are only three estimated parameters (D is inferred by the three haplotype frequencies) and 3 df. It is important to note that in the latter case neither cubic solution nor iteration is necessary as the haplotype frequencies can be directly counted from the diplotype data. If the user believes that there are only three alleles and hence six diplotypes, but there are nonzero values for any of the other three possible diplotypes, then reconsideration of the technical veracity of the data and of the homogeneity of the population sample would be wise.

Results

Solutions are considered biologically possible when \hat{f}_{11} and the derived \hat{f}_{12} , \hat{f}_{21} and \hat{f}_{22} all fall within the range 0 to 1 (i.e. \hat{f}_{11} , \hat{f}_{12} , \hat{f}_{21} , $\hat{f}_{22} \in [0,1]$) and add up to 1. This constraint is tighter than those described elsewhere [22] as it relies on the inherent assumption of representative sampling and HWE, an extreme chance distortion of which could lead to three solutions at SNP allele frequencies of 0.5 in sample data drawn from a population (if all samples are heterozygous at both loci the following are possible: all could be diplotype 11/22, all could be diplotype 12/21, or there could be a combination of both).

Number of solutions to the cubic equation with simulated data

We have calculated the number of possible solutions to the cubic equation for genotypes of simulated pairs of SNPs with a range of allele frequencies for a range of sample sizes. The genotype numbers were calculated assuming HWE with a wide range of LD situations for the two SNPs. This was achieved by simulating all combinations of haplotype frequencies between 0 and 1, at intervals of 1/55, that add up to 1. These haplotype frequencies were then converted to diplotype frequencies according to Hardy-Weinberg equilibrium. The results are plotted in Figure 1. Small samples result in minor deviations from sample HWE, allowing more than one solution. The smaller the sample size, the greater the range of allele frequencies over which this occurs. A sample of 10 subjects allows more than one biologically possible solution at a wide range of allele frequencies (Figure 1A). With 60 individuals a broad range of allele frequencies is still affected (Figure 1B) - this has implications for analyses based on the HapMap CEU dataset of 60 unrelated individuals [23,24]. At 100 individuals (Figure 1C) the problem is limited to allele frequencies below 15% (Figure 1C), while the plot for 1000 individuals shows no condition under which there is more than one biologically possible solution (Figure 1D). This last observation is because under perfect sample HWE (infinite samples) the number of biologically possible solutions is always 1, despite the number of real solutions exceeding 1 at lower allele frequencies (data not shown).

Number of solutions to the cubic equation with real data

We have also calculated the number of solutions to equation 1 for a set of real data from the HapMap project [23,24]. These were a selection of SNPs from the *ACE-GH1* region of chromosome 17 for the CEU population (60 unrelated individuals). Figure 2A shows that at the lower allele frequencies the possibility of more than one real solution to the cubic equation begins to arise. This is

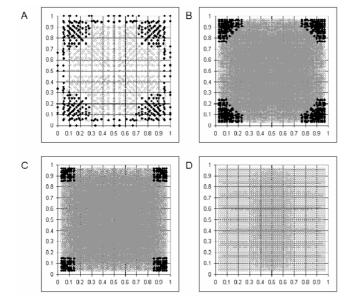


Figure I
Simulated data in which HWE is observed to the limit of rounding errors (whole number values for counts of individuals). (A) Number of biologically possible solutions to the cubic equation in (A) 10 individuals; (B) 60 individuals; (C) 100 individuals (D) 1000 individuals. x-axis: allele frequency of SNP1, y-axis: allele frequency of SNP2. Black = more than one solution. Grey = one solution.

consistent with the simulated data for 60 samples (Figure 1B), except that a broader range of allele frequencies is affected. This is probably due to the inherent errors of real data increasing the deviations from HWE relative to nearperfect simulated data. In most cases of multiple solutions only two of the three real roots are biologically possible. Figure 2B compares these two values, indicating that in most cases the differences in estimated haplotype are small. In the minority of cases with three solutions these fit the same pattern. However, this can have major consequences for the calculation of D' (as illustrated in Figure 3). Note that D' and r² behave quite differently in this respect, and r² is much less affected. However, as a |D'| of 1 indicates the existence of three or less haplotypes (r² of 1 indicates two haplotypes), |D'| is a good indicator of haplotype block structure, with a value of exactly 1 suggesting little or no recombination between two loci, and a value less than 1 supporting a break-down of LD. In fact CubeX provides both D' and r2, allowing the user to select their measure of preference. Figure 4 illustrates the relationship between these two measures in the simulated and real datasets, which clarifies how a large |D'| value can be observed with a low r² value, but the key point is that a |D'| of 1 indicates complete LD (i.e. three or less haplotypes) despite a low r2.

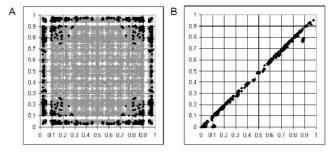


Figure 2
Evaluation of number of solutions for real data. (A)
Number of biologically possible solutions over a range of
allele frequencies using a large sample of SNP data (Chr.
17:60 to 60.5 MB, 121 SNPs) from the HapMap project
[23,24]. x-axis: allele frequency of SNP1, y-axis: allele frequency of SNP2. Black = more than one solution. Grey = one
solution. (B) Comparison of two solutions within the dataset.
x-axis: higher value solution, y-axis: lower value solution.

Comparison of the cubic exact solution with other approaches

For the purposes of comparison we have analysed two datasets with PHASE [16], MIDAS [7] (Hill EM) and CubeX. The first is a dataset of directly haplotyped samples comprising 80 subjects from 3 ethnic groups (Asian, African and Caucasian) for APOE [25]. Although all but one SNP was in Hardy-Weinberg equilibrium, this dataset has the potential to invalidate some of the assumptions of the programs due to the mixture of ethnicity. However, this provides a useful substrate on which to test the influence of stratification on the outcome of the cubic exact solution. The second dataset is a set of multi-locus phased data from HapMap CEU samples [23,24] for the IGF2 gene region. Although these have not been directly haplotyped, the multi-locus phased haplotypes are expected to be very accurate, and this dataset comprises Caucasians, so will not suffer from the same stratification issues. We tested the programs on pair-wise subsets of these data.

For the *APOE* [25] dataset the data are presented in Additional File 1, with a selected summary in Table 1. The subset in Table 1 demonstrate the advantage of being provided with all possible solutions by CubeX, but also demonstrates that all three approaches can be wrong. To summarise the outcome, PHASE [16] and MIDAS [7] (Hill EM) both matched the real counts in 28 of 36 SNP pairs, while CubeX matched real counts in 33 of 36 SNP pairs (for one of its solutions). However, in five of those cases the user would need to determine which of the two CubeX solutions to use based on their prior knowledge of the LD structure in the region (i.e. do they expect three or four haplotypes). This comparison confirms the risk of EM finding a local maximum when there is more than one biologically possible solution, and suggests that CubeX

Results

For an explanation of the analysis and results please see noise below Number of biologically possible solutions: 2. B is the most likely solution $See X^{*}$ table below 3x3

| | H | ap lo typ e t | freq ue nci | D' statistics | | | | | |
|----------|--------|---------------|-------------|---------------|-------|----------------|-------|--|--|
| Solution | f11 | f12 | f21 | f22 | D, | r ² | X^2 | | |
| β | 0.1667 | 0.5833 | 0.0167 | 0.2333 | 0.636 | 0.0303 | 1.82 | | |
| Y | 0.1833 | 0.5667 | -0.0 | 0.25 | 1.0 | 0.0748 | 4.49 | | |

3x3 table of observed and expected diplotype numbers

Black numbers on white are original data entered Coloured numbers on coloured background represent the solutions from the table above

| | | SNP 2 | | | | | | | |
|-------|----|-------|------|------|--|--|--|--|--|
| | | 11 | 12 | 22 | | | | | |
| | | 1 | 10 | 22 | | | | | |
| | 11 | 1.7 | 11.7 | 20.4 | | | | | |
| | | 2.0 | 12.5 | 19.3 | | | | | |
| | 12 | 0 | 10 | 14 | | | | | |
| SNP 1 | | 0.3 | 5.8 | 16.3 | | | | | |
| | | -0.0 | 5.5 | 17.0 | | | | | |
| | | 0 | 0 | 3 | | | | | |
| | 22 | 0.0 | 0.5 | 3.3 | | | | | |
| | | 0.0 | -0.0 | 3.8 | | | | | |

Solution Z of 3x3 5.7496

This is a K^2 (5 degrees of freedom) of the 3x3 table. The higher the value, the less good the fit of the observed haplotypes to Harv-Weinberg equilibrium. A value greater than 11.07 indicates a significance p=0.05. If there are two or more solutions, the lower values are more likely. However, a significant value indicates genotype data out of Hardy-Weinberg equilibrium, a problem that should be addressed before interpreting these results.

Minimum biologically possible f11:0.1 Maximum biologically possible f11:0.18333

Number of impossible solutions: 1 α : $f_{11} = 0.225$

SNP 1 allele 1 frequency = 0.75 SNP 2 allele 1 frequency = 0.183

- Here f₁₁ refers to an estimated haplotype frequency for allele 1 at locus 1 and allele 1 at locus 2 (likewise for other haplotypes 12,21 and 22). The character f should have a "hat" to indicate that it is estimated but 1 that limitations prevent this. f₁₁ is based on direct solution of the cubic equation expressing the phase uncertain double heteroxygotes (middle square of the 3x3) and overall model in terms of estimated f₁₁
- - 1. random mating 2. Hardy-Weinberg equilibrium at both loci
- Expectation-maximisation algorithms rely on an iteration rather than direct solution. The direct solution will display both the most likely (which EM should ideally reach) and other
- direct solution will display both the most likely (which EM should ideally reach) and possible solutions where iterations may converge in error. At values represent difference between observed and expected diplotype frequencies. In perfect Hardy-We inberge equilibrium "expected" numbers will be whole numbers matching the "observed" diplotype numbers. However, imperfect Hardy-Weinberg proportions will result in impossible "fractions" of individuals, which are shown for comparison to observations.

Return to input

Figure 3

Screenshot of results screen from CubeX online analysis program. In this example there are two biologically possible solutions. Results for both are shown (upper table), and observed (input values) and expected diplotype frequencies (for the two solutions) displayed for comparison (lower table).

may offer advantages in stratified datasets or datasets with low SNP minor allele frequencies (confirming the results from simulated data above).

For the HapMap [23,24] IGF2 region data (comprising SNPs rs3802971, rs734351, rs3213221, rs4244808, rs1003483, rs3741208, rs1004446, rs4320932 and

rs7924316) CubeX gives only one solution in all cases, and there is little difference between the outcome of the three approaches (Additional File 2). This confirms that in situations of higher allele frequencies there is less of an issue with multiple biologically possible solutions to the cubic equation, and iterative approaches are completely acceptable.

Discussion

We have written an online program, "CubeX", to enable simple analysis of the biologically possible estimated haplotypes for pairs of biallelic markers. This program takes data from a pair of markers as a standard 3 × 3 table of nine diplotypes, generates cubic exact solutions to equation 1 and generates output in the format shown in Figure 3. The number of possible solutions is shown, followed by haplotype frequencies and LD statistics for those solutions. Below that a duplicate of the 3×3 input table is displayed with the addition of expected absolute diplotype frequencies calculated from the haplotype frequencies. The difference between these and the input data are subjected to a χ^2 test, which effectively tests sample deviation from the null hypothesis of HWE for the diplotypes formed of the four haplotypes. However, the interpretation of solutions depends on the prior hypothesis. In the example in Figure 3, although solution γ exhibits a slightly worse χ^2 fit than solution β_t the former is consistent with a prior hypothesis of only three of the four haplotypes existing (see Figure 5 in reference [7]), which is biologically likely in the absence of recombination between any two loci. In fact, in all tested cases in Figure 2 generating more than one solution, the diplotype data included zero values in at least one corner cell and the two adjacent edge cells of the 3×3 (i.e. where one possible solution has a |D'| = 1, although it should be noted that more than one solution can occur without zero values if double heterozygotes are greatly over-represented). This suggests that the principal issue is whether three or four haplotypes exist, and in these cases the prior hypothesis (based on distance and recombination rates) is of utmost importance. If input data for individual SNPs are significantly out of HWE a warning message is given at the top of the page. For completeness, the biologically impossible real number solutions are displayed at the bottom, along with minimum and maximum biologically possible values for

 f_{11} and allele frequencies. This program provides a convenient utility for researchers to both analyse data for haplotype frequencies and LD statistics and to check previous analyses for potential problems caused by multiple solutions.

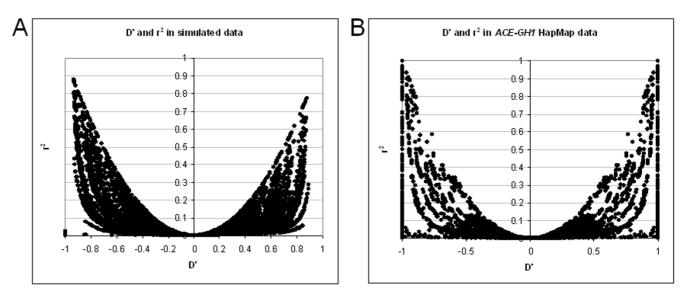


Figure 4
The range of LD in datasets using the CubeX tool to calculate r2 and D'. (A) Simulated data. D' on x-axis, r² on y axis. (B) Real SNP data (Chr. 17:60 to 60.5 MB, 121 SNPs) from the HapMap project [23,24]. D' on x-axis, r² on y axis.

Under perfect sample HWE the frequencies of all haplotypes can be directly inferred from the corresponding corner diplotypes of the 3 × 3. For example: $n_{11} = n \, \hat{f}_{11}^2$, so $\hat{f}_{11} = \sqrt{\frac{n_{11}}{n}}$. That being the case there are only two possi-

ble values for \hat{f}_{11} , one positive and one negative, the latter being biologically impossible. Perfect sample HWE therefore results in only a single biologically possible solution to the cubic equation. In the case of extreme sample HWD where all samples fall within the middle cell of the 3 × 3,

 f_{11} can contribute either a half, a quarter or none of the haplotypes to the middle cell. There are therefore three biologically possible solutions under conditions of extreme sample HWD. The results from real data confirm that in some cases more than one biologically possible solution to the cubic equation for haplotype frequency can exist. The simulations suggest that this occurs where small sample size, sampling errors or non-random mating result in a distortion of sample HWE, and demonstrates the importance of testing HWE before haplotype analyses. The greater the distortion of sample HWE the higher the allele frequency at which more than one solution can occur (hence, as described above, three solutions can occur at allele frequencies of 0.5 if all samples are heterozygous at both loci). In these cases the cubic exact algorithm gives all possible solutions and a test of HWE, while an iteration-based method would only give one. This supports the hypothesis that the cubic exact approach is supe-

rior to iteration-based methods in real-world datasets where sample data rarely fit exactly to HWE (note that sample may differ from population in HWE statistics here we refer to sample HWE). This is particularly important in the analysis of low frequency SNPs and paucimorphisms [26-28], for which different solutions can significantly distort D' results, despite the relatively similar solutions giving similar r² results. In all the observed data with two solutions there were no occasions in which r² exceeded 0.3 for any biologically possible solution, and in most cases there is only a small difference in r² between biologically possible solutions. The largest effect is on D'. On the basis of empirical data and using different approaches to inference Wong et al showed that coding SNPs with minor allele frequencies <0.06 are likely to be of functional importance [29], and rarer alleles, haplotypes and diplotypes of causal importance have emerged in numerous disease contexts (eg. inflammatory bowel disease, hemochromatosis). In addition to being applicable and giving exact evaluation for D' analysis of common SNPs, the cubic exact solution may prove of particular value for evaluating "post-HapMap" and "post-dbSNP" rarer haplotypes, for fully evaluating D' estimates from datasets with greater deviations from the random mating and HWE assumptions and for fully evaluating LD in small datasets.

Finally, we have demonstrated by comparison with PHASE [16] and MIDAS [7] (Hill EM) that in certain situ-

Table I: Illustrative examples of comparison of CubeX with PHASE [16] and MIDAS [7] (Hill EM).

| | | | Haplotype frequencies (rounded to 5 decimal places) | | | | | | Haplotype numbers (rounded to nearest haplotype) | | | | | | |
|---------|----------|-----------|---|--------------------|--------------------|----------------|---------------|----------------|--|-----------------|-----------------|----------------|---------------|-------|--|
| Example | SNP pair | Haplotype | REAL frequency | PHASE frequency | MIDAS frequency | CUBEX alpha | CUBEX beta | CUBEX gamma | REAL number | PHASE number | MIDAS number | CUBEX alpha | CUBEX beta | CUBEX | |
| I | Pair I_2 | AC | 0.0875 | 0.08689 | 0.0875 | na | 0.0875 | na | 14 | 14 | 14 | na | 14 | na | |
| | Pair I_2 | AT | 0.725 | 0.72561 | 0.725 | na | 0.725 | na | 116 | 116 | 116 | na | 116 | na | |
| | Pair I_2 | TC | 0 | 0.00061 | 0 | na | 0 | na | 0 | 0 | 0 | na | 0 | na | |
| | Pair1_2 | TT | 0.1875 | 0.18689 | 0.1875 | na | 0.1875 | na | 30 | 30 | 30 | na | 30 | na | |
| 2 | Pairl_5 | AG | 0.75 | 0.75318 | 0.75478 | 0.75478 | na | 0.75 | 120 | 121* | 121 * | 121 * | na | 120 | |
| | Pair1_5 | AA | 0.0625 | 0.05932 | 0.05772 | 0.05772 | na | 0.0625 | 10 | 9 * | 9 * | 9* | na | 10 | |
| | Pair1_5 | TG | 0.1875 | 0.18432 | 0.18272 | 0.18272 | na | 0.1875 | 30 | 29 * | 29 * | 29 * | na | 30 | |
| | Pair1_5 | TA | 0 | 0.00318 | 0.00478 | 0.00478 | na | 0 | 0 | I * | I * | I * | na | 0 | |
| 3 | Pair1_9 | АТ | 0.05625 | 0.06477 | 0.05633 | na | 0.0563 | 0.075 | 9 | 10* | 9 | na | 9 | 12* | |
| | Pair1_9 | AC | 0.75625 | 0.74773 | 0.75617 | na | 0.7562 | 0.7375 | 121 | 120 * | 121 | na | 121 | 118* | |
| | Pair I_9 | TT | 0.01875 | 0.01023 | 0.01867 | na | 0.0187 | 0 | 3 | 2 * | 3 | na | 3 | 0 * | |
| | Pair1_9 | TC | 0.16875 | 0.17727 | 0.16883 | na | 0.1688 | 0.1875 | 27 | 28 * | 27 | na | 27 | 30 * | |
| 4 | Pair2_3 | CG | 0.05625 | 0.04724 | 0.0465 | 0.0465 | na | na | 9 | 8 * | 7 * | 7* | na | na | |
| | Pair2_3 | CT | 0.03125 | 0.04026 | 0.041 | 0.041 | na | na | 5 | 6* | 7 * | 7* | na | na | |
| | Pair2_3 | TG | 0.48125 | 0.49026 | 0.491 | 0.491 | na | na | 77 | 78 * | 79 * | 79 * | na | na | |
| | Pair2_3 | TT | 0.43125 | 0.42224 | 0.4215 | 0.4215 | na | na | 69 | 68 * | 67 * | 67 * | na | na | |
| 5 | pair5_9 | GT | 0.075 | 0.07313 | 0.06664 | na | 0.0666 | 0.075 | 12 | 12 | 11 * | na | 11* | 12 | |
| | pair5_9 | GC | 0.8625 | 0.86437 | 0.87086 | na | 0.8709 | 0.8625 | 138 | 138 | 139 * | na | 139 * | 138 | |
| | pair5_9 | AT | 0 | 0.00187 | 0.00836 | na | 0.0084 | 0 | 0 | 0 | 1* | na | 1* | 0 | |
| | pair5_9 | AC | 0.0625 | 0.06063 | 0.05414 | na | 0.0541 | 0.0625 | 10 | 10 | 9 * | na | 9 * | 10 | |

Page 7 of 9 (page number not for citation purposes)

A selection of comparisons using direct haplotyped APOE data [25]. Full data are present as a additional table. For haplotype numbers (rounded to the nearest number) incorrect answers are marked **, correct answers are unmarked. Examples: (I) Phase, MIDAS and CubeX (I solution) give correct answer. (2) Only CubeX gives the correct answer as one of its two solutions. (3) MIDAS and CubeX give the correct answer, PHASE and the other CubeX solution are wrong. (4) All three approaches are wrong. (5) PHASE and CubeX give the correct answer, MIDAS and the other CubeX solution are wrong.

ations (low minor allele frequency, population stratification) the cubic exact approach can perform better for pairwise analyses than alternative approaches by indicating the existence of multiple solutions. However, our findings confirm that in most other situations iterative approaches are robust and accurate.

Conclusion

We present a comprehensive analysis of the consequences of different variables on the number of solutions to the cubic equation for haplotype frequency. Our analyses demonstrate that lower allele frequencies, lower sample numbers and a possible |D'| value of 1 can result in more than one solution. This has significant implications for the calculation of LD in small sample sizes and with rarer alleles that may have particular disease relevance. This evaluation provides essential information for an understanding of the limitations of LD estimation, which is particularly relevant for genome-wide analyses (where sample sizes and allele frequencies can be low). Finally, we present a program "CubeX", freely available as an online program, which provides each of the biologically possible cubic exact solution(s) to equation 1 for haplotype frequency, enabling the user to identify the solution that best fits their prior hypothesis for number of haplo-

Availability and Requirements

Project name: CubeX

Project home page: http://www.oege.org/software/cubex

Operating system(s): Platform independent (web-based)

Programming language: Python http://www.python.org

Licence: CubeX licence available from http://www.oege.org/software/cubex

Any restrictions to use by non-academics: royalty-free use allowed within terms of licence

Abbreviations

EM - Expectation-Maximisation

HWE - Hardy-Weinberg Equilibrium

LD - Linkage Disequilibrium

Authors' contributions

TRG wrote the CubeX program, ran the simulations and analyses and drafted the manuscript. SR advised on LD calculation and output format, tested the program and contributed to the manuscript. INMD drafted the solution to the cubic equation, advised on methods, tested the pro-

gram and contributed to the manuscript. All authors read and approved the final manuscript.

Additional material

Additional file 1

Comparisons of PHASE, MIDAS and CubeX on APOE data (from [25]). A comparison of PHASE, MIDAS and CubeX for pairwise analysis of genotype data derived from directly observed multi-locus haplotypes. Click here for file

[http://www.biomedcentral.com/content/supplementary/1471-2105-8-428-S1.pdf]

Additional file 2

Comparisons of PHASE, MIDAS and CubeX on HapMap IGF2 region data (from http://www.hapmap.org. [23,24]). A comparison of PHASE, MIDAS and CubeX for pairwise analysis of genotype data derived from statistically inferred long-range multi-locus haplotypes.

Click here for file

[http://www.biomedcentral.com/content/supplementary/1471-2105-8-428-S2.pdf]

Acknowledgements

TRG is funded by a BHF (British Heart Foundation) Intermediate Fellowship (FS/05/065/19497), SR by a HOPE (Wessex Medical Trust) fellowship and work in our laboratory by the Medical Research Council (UK) (Programme Grant G9800748). We thank an anonymous reviewer for their suggestion of a comparison with PHASE on the APOE dataset [25].

References

- Weiss KM, Clark AG: Linkage disequilibrium and the mapping of complex human traits. Trends in Genetics 2002, 18:19-24.
- Palmer LJ, Cardon LR: Shaking the tree: mapping complex disease genes with linkage disequilibrium. The Lancet 2005, 366:1223-1234.
- Ardlie KG, Kruglyak L, Seielstad M: Patterns of linkage disequilibrium in the human genome. Nat Rev Genet 2002, 3:299-309.
- Hill WG: Estimation of linkage disequilibrium in randomly mating populations. Heredity 1974, 33:229-239.
- Abecasis GR, Cookson WO: GOLD--graphical overview of linkage disequilibrium. Bioinformatics 2000, 16:182-183.
- Pettersson F, Jonsson O, Cardon LR: GOLDsurfer: three dimensional display of linkage disequilibrium. Bioinformatics 2004, 20:3241-3243.
- Gaunt TR, Rodriguez S, Zapata C, Day IN: MIDAS: software for analysis and visualisation of interallelic disequilibrium between multiallelic markers. BMC Bioinformatics 2006, 7:227.
- Barrett JC, Fry B, Maller J, Daly MJ: Haploview: analysis and visualization of LD and haplotype maps. Bioinformatics 2005, 21:263-265.
- Jorde LB: Linkage disequilibrium and the search for complex disease genes. Genome Res 2000, 10:1435-1444.
- Mueller JC: Linkage disequilibrium for different scales and applications. Brief Bioinform 2004, 5:355-364.
- 11. Weale ME: A survey of current software for haplotype phase inference. Hum Genomics 2004, 1:141-144.
- Salem RM, Wessel J, Schork NJ: A comprehensive literature review of haplotyping software and methods for use with unrelated individuals. Human Genomics 2005, 2:39-66.
- 13. Wang L, Xu Y: Haplotype inference by maximum parsimony. *Bioinformatics* 2003, 19:1773-1780.
- Stephens M, Scheet P: Accounting for decay of linkage disequilibrium in haplotype inference and missing-data imputation. Am J Hum Genet 2005, 76:449-462.

- Stephens M, Donnelly P: A comparison of bayesian methods for haplotype reconstruction from population genotype data. Am J Hum Genet 2003, 73:1162-1169.
- Stephens M, Smith NJ, Donnelly P: A new statistical method for haplotype reconstruction from population data. Am J Hum Genet 2001, 68:978-989.
- Nickalls RWD: A new approach to solving the cubic: Cardan's solution revealed. The Mathematical Gazette 1993, 77:354-359.
- Luo ZW, Suhai S: Estimating Linkage Disequilibrium Between a Polymorphic Marker Locus and a Trait Locus in Natural Populations. Genetics 1999, 151:359-371.
- Foundation PS: The Python Programming Language. 2006 [http://www.python.org].
- Lewontin RC: The interaction of selection and linkage. I. General considerations; heterotic models. Genetics 1964, 49:49-67.
- Hill WG, Robertson A: Linkage disequilibrium in finite populations. Theor Appl Genet 1968:135-156.
- Mano S, Yasuda N, Katoh T, Tounai K, Inoko H, Imanishi T, Tamiya G, Gojobori T: Notes on the Maximum Likelihood Estimation of Haplotype Frequencies. Annals of Human Genetics 2004, 68:257-264
- 23. The International HapMap Project. Nature 2003, 426:789-796.
- Consortium TIHM: A haplotype map of the human genome. Nature 2005, 437:1299-1320.
- Orzack SH, Gusfield D, Olson J, Nesbitt S, Subrahmanyan L, Stanton VP Jr.: Analysis and Exploration of the Use of Rule-Based Algorithms and Consensus Methods for the Inferral of Haplotypes. Genetics 2003, 165:915-928.
- Day INM, Alharbi KK, Smith MJ, Aldahmesh MA, Chen X, Lotery AJ, Pante-de-Sousa G, Hou G, Ye S, Eccles DM, Cross NCP, Fox KR, Rodriguez S: Paucimorphic Alleles versus Polymorphic Alleles and Rare Mutations in Disease Causation: Theory, Observation and Detection. Current Genomics 2004, 5:431-438.
- 27. Alharbi KK, Aldahmesh MA, Spanakis E, Haddad L, Whittall RA, Chen X, Rassoulian H, Smith MJ, Sillibourne J, Ball NJ, Graham NJ, Briggs PJ, Simpson IA, Phillips DIW, Lawlor DA, Ye S, Humphries SE, Cooper C, Smith GD, Ebrahim S, Eccles DM, Day INM: Mutation scanning by meltMADGE: Validations using BRCAI and LDLR, and demonstration of the potential to identify severe, moderate, silent, rare, and paucimorphic mutations in the general population. Genome Res 2005, 15:967-977.
- Alharbi KK, Spanakis E, Tan K, Smith MJ, Aldahmesh MA, O'Dell SD, Sayer AA, Lawlor DA, Ebrahim S, Davey Smith G, O'Rahilly S, Farooqi S, Cooper C, Phillips DI, Day IN: Prevalence and functionality of paucimorphic and private MC4R mutations in a large, unselected European British population, scanned by melt-MADGE. Hum Mutat 2007, 28(3):294-302.
- Wong GKS, Yang Z, Passey DA, Kibukawa M, Paddock M, Liu CR, Bolund L, Yu J: A Population Threshold for Functional Polymorphisms. Genome Res 2003, 13:1873-1879.

Publish with **Bio Med Central** and every scientist can read your work free of charge

"BioMed Central will be the most significant development for disseminating the results of biomedical research in our lifetime."

Sir Paul Nurse, Cancer Research UK

Your research papers will be:

- available free of charge to the entire biomedical community
- peer reviewed and published immediately upon acceptance
- cited in PubMed and archived on PubMed Central
- yours you keep the copyright

Submit your manuscript here: http://www.biomedcentral.com/info/publishing_adv.asp

